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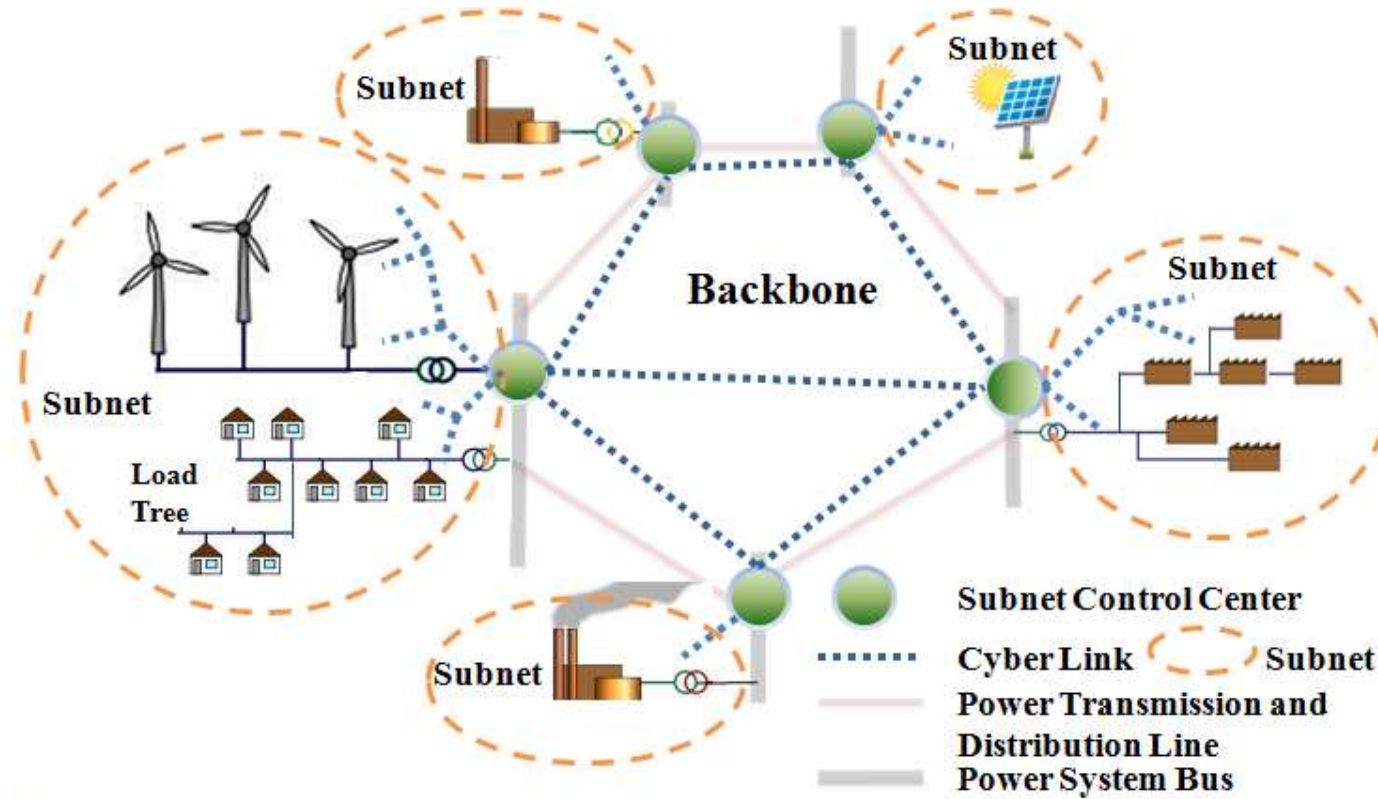
# Distributed Estimation and Detection for Smart Grid

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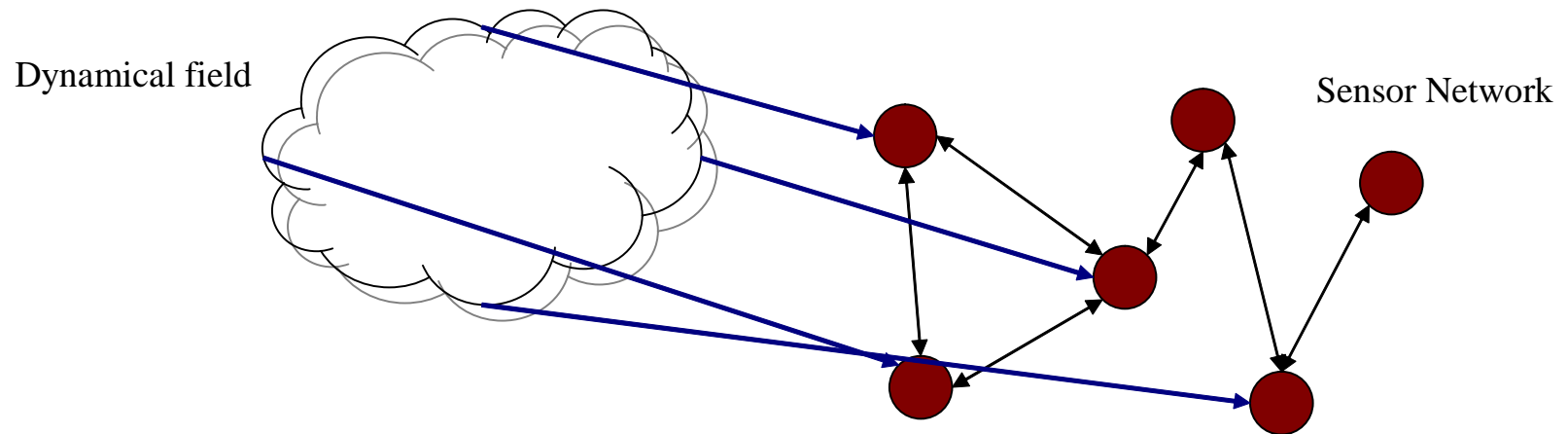
- Joint Work with: S. Kar (CMU), R. Tandon (Princeton), H. V. Poor (Princeton), and J. M. F. Moura (CMU)

# Distributed Estimation/Detection in Smart Grid



Estimation: Large-scale power system state  
Detection: Attack or anomaly in the system

# Distributed Estimation over Linear Dynamical Systems



Field/Signal Process:  $\{\mathbf{x}_k\} \in \mathbb{R}^M$

$$\mathbf{x}_{(k+1)\Delta} = \mathcal{F}\mathbf{x}_{k\Delta} + \mathbf{w}_k$$

Local observation at sensor  $n$ :  $\{\mathbf{y}_{k\Delta}^n\} \in \mathbb{R}^{m_n}$

$$\mathbf{y}_{k\Delta}^n = \mathcal{C}_n \mathbf{x}_{k\Delta} + \mathbf{v}_k^n$$

- Applications: Smart Grid, Social Network, etc...

## Assumptions on the Signal Model

We impose the following global assumptions on the signal/observation model:

- **Stabilizability-Assumption (S.1):** The pair  $(\mathcal{F}, Q^{1/2})$  is stabilizable. The non-degeneracy (positive definiteness) of  $Q$  ensures this.
- **Global Detectability-Assumption (D.1):** The pair  $(\mathcal{C}, \mathcal{F})$  is detectable, where  $\mathcal{C} = [\mathcal{C}_1^T \cdots \mathcal{C}_N^T]^T$ .

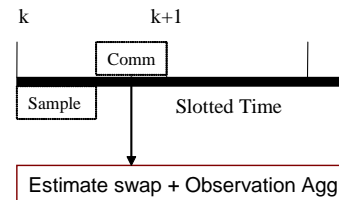
**Remark 1.** *We do not assume local detectability at each agent. These assumptions are required even by the centralized estimator to achieve a stable estimation error. We later show that under rather weak conditions on the inter-sensor communication, the global detectability assumption is also sufficient for our distributed scheme to achieve stable estimation errors at each sensor.*

## Distributed Filtering Architecture

- The agents (sensors) need to collaborate for *successful* estimation of  $\{\mathbf{x}_k\}$ .
- Basic Sensing and Communication Architecture:
  - Inter-agent communication rate is measured as the link activation rate
  - Inter-agent communication is assumed perfect
- Communication is constrained by the network topology; and only  $\bar{\gamma} > 0$  rounds of inter-agent message passing may be allowed per epoch  $(k\Delta, (k+1)\Delta]$  on an average.

## M-GIKF Scheme

The overall communication rate  $\bar{\gamma}$  is split over two phases:



- **Estimate Swapping:**

- Agents perform pairwise swapping of previous epoch's estimates (if the corresponding link is active)
- *Estimate Swapping guarantees asymptotic stability of local estimation errors under weak global detectability assumptions*

- **Observation Aggregation:**

- Agents use the remaining communication to disseminate their instantaneous observations to the neighbors
- *Observation Aggregation improves estimation performance*

## Communication Rate Constraint

### Define:

- $M^e(k)$ : Number of sensor communications at the  $k$ -th epoch for estimate swapping
- $M^o(k)$ : Number of sensor communications at the  $k$ -th epoch for observation aggregation

**Communication Constraint:** Since  $\bar{\gamma} > 0$  is given, we require:

$$\limsup_{k \rightarrow \infty} \frac{1}{k} \sum_{j=0}^{k-1} (M^e(k) + M^o(k)) \leq \bar{\gamma} \text{ a.s.}$$

## GS Protocol: Estimate Swapping

**Connectivity - Assumption (C.1):** The maximal graph  $(V, \mathcal{E})$  is connected, where  $V$  denotes the set of  $N$  agents and  $\mathcal{E}$  the set of allowable links.

**Proposition 1.** *Under (C.1) we have the following:*

- *The sequence  $\{A^e(k)\}$  of estimate swapping adjacency matrices is independent and identically drawn from a special distribution  $\mathcal{D}$*
- *The mean adjacency matrix  $\bar{A}^e$  is irreducible and aperiodic*
- *The average number of inter-sensor communications due to estimate swapping satisfies almost surely (a.s.)*

$$\bar{M}^e = \lim_{k \rightarrow \infty} (1/k) \sum_{i=0}^{k-1} M^e(k) < \bar{\gamma}/2$$



## GS Protocol: Observation Aggregation

Adjacency matrix sequence drawn following a Poisson process of rate  $\bar{\gamma}/2$

**Proposition 2.** *Under (C.1) we have the following:*

- *The average number of inter-agent communications due to observation aggregation per epoch is  $\bar{\gamma}/2$ .*
- *The sequence  $\{\mathcal{I}_k^n\}$  denotes the set of observations (w.r.t. node indices) available at sensor  $n$  at the end of the epochs; then for every epoch  $k$  and every sensor  $n$ ,*

$$\mathbb{P}(\mathcal{I}_k^n = [1, \dots, N]) > 0$$

## Reasonable Communication Scheme

**Definition 1.** *A communication scheme (for estimate swapping and observation aggregation) is said to be reasonable if*

- (E.1) *The estimate swapping adjacency matrices  $\{A^e(k)\}$  are i.i.d. The mean matrix  $\bar{A}$  is doubly stochastic, irreducible, and aperiodic.*
- (E.2) *The sequences  $\{\mathcal{I}_k^n\}$  are i.i.d. for each  $n$ , independent of the estimate swapping and satisfy  $\mathbb{P}(\mathcal{I}_k^n = [1, \dots, N]) > 0$ .*
- (E.3) *The average number of inter-sensor communications (including both the estimate swapping and observation aggregation steps) per epoch is less than or equal to  $\bar{\gamma}$ , where  $\bar{\gamma} > 0$  is a predefined upper bound on the communication rate.*

**Remark 2.** *The previous GS protocol is a reasonable scheme under (C.1).*

## M-GIKF: Estimate Update

### Define:

- $(\hat{\mathbf{x}}_{k|k-1}^n, \hat{P}_k^n)$ : Estimate (state) at sensor  $n$  of  $\mathbf{x}_k$  based on *information* till time  $k - 1$
- $n_k^{\rightarrow}$ : Neighbor of sensor  $n$  at time  $k$  w.r.t.  $A^e(k)$ .

With *estimate swapping*: swap  $(\hat{\mathbf{x}}_{k|k-1}^n, \hat{P}_k^n)$  and  $(\hat{\mathbf{x}}_{k|k-1}^{n_k^{\rightarrow}}, \hat{P}_k^{n_k^{\rightarrow}})$ ;

With *observation aggregation*: sensor  $n$  collects:  $\mathbf{y}_k^{\mathcal{I}_n^k}$

### Update Rule:

$$\hat{\mathbf{x}}_{k+1|k}^n = \mathbb{E} \left[ \mathbf{x}_{k+1} \mid \hat{\mathbf{x}}_{k|k-1}^{n_k^{\rightarrow}}, \hat{P}_k^{n_k^{\rightarrow}}, \mathbf{y}_k^{\mathcal{I}_n^k} \right]$$

$$\hat{P}_{k+1}^n = \mathbb{E} \left[ \left( \mathbf{x}_{k+1} - \hat{\mathbf{x}}_{k+1|k}^n \right) \left( \mathbf{x}_{k+1} - \hat{\mathbf{x}}_{k+1|k}^n \right)^T \mid \hat{\mathbf{x}}_{k|k-1}^{n_k^{\rightarrow}}, \hat{P}_k^{n_k^{\rightarrow}}, \mathbf{y}_k^{\mathcal{I}_n^k} \right]$$

## M-GIKF: Estimate Update (*Contd.*)

- The filtering steps can be implemented through time-varying Kalman filter recursions
- The sequence  $\{\hat{P}_k^n\}$  of conditional predicted error covariance matrices at sensor  $n$  satisfies the random Riccati recursion:

$$\hat{P}_{k+1}^n = \mathcal{F} \hat{P}_k^{n_{\vec{k}}} \mathcal{F}^T + \mathcal{Q} - \mathcal{F} \hat{P}_k^{n_{\vec{k}}} \mathcal{C}_n^T \left( \mathcal{C}_n \hat{P}_k^{n_{\vec{k}}} \mathcal{C}_n^T + \mathcal{R}_n \right)^{-1} \mathcal{C}_n \hat{P}_k^{n_{\vec{k}}} \mathcal{F}^T$$

- The sequence  $\{\hat{P}_k^n\}$  is random due to the random neighborhood selection functions  $n_{\vec{k}}$  and  $\mathcal{I}_n^k$
- The goal is to study asymptotic properties of  $\{\hat{P}_k^n\}$  at every sensor  $n$ 
  - In what sense  $\{\hat{P}_k^n\}$  is stable
  - In what sense they reach *agreement*

## Switched Riccati Iterates

Let  $\mathfrak{P}$  denote a generic subset of  $[1, \dots, N]$ .

**Define:**

- $\mathcal{C}_{\mathfrak{P}}$ : The stack of  $\mathcal{C}_j$ 's for all  $j \in \mathfrak{P}$
- $f_{\mathfrak{P}}(\cdot)$ : The Riccati operator given by

$$f_{\mathfrak{P}}(X) = \mathcal{F}X\mathcal{F}^T + \mathcal{Q} - \mathcal{F}X\mathcal{C}_{\mathfrak{P}}^T (\mathcal{C}_{\mathfrak{P}}X\mathcal{C}_{\mathfrak{P}}^T + \mathcal{R}_n)^{-1} \mathcal{C}_{\mathfrak{P}}X\mathcal{F}^T$$

*The conditional prediction error covariances are then updated as:*

$$\hat{P}_n(k+1) = f_{\mathcal{I}_k^n} \left( \hat{P}_{n_k^{\rightarrow}}(k) \right), \quad \forall n$$

**Remark 3.** *The sequence  $\{\hat{P}_n(k)\}$  evolves as a random Riccati iterate with non-stationary switching.*

## M-GIKF: Main Results on Convergence

**Theorem 1.** *Consider the M-GIKF under assumptions (S.1), (D.1), (C.1), and (E.1)-(E.3). Then,*

- (1) *Let  $q$  be a uniformly distributed random variable on  $[1, \dots, N]$  independent of the sequences  $\{A^e(k)\}$  and  $\{\mathcal{I}_n^k\}$ . Then, the sequence  $\{\hat{P}_q(k)\}$  converges weakly to an invariant distribution  $\mu^{\bar{\gamma}}$  on  $\mathbb{S}_+^N$ , i.e.,*

$$\hat{P}_q(k) \implies \mu^{\bar{\gamma}}.$$

- (2) *The performance approaches the centralized one exponentially over  $\bar{\gamma}$ .*
- (3) *For each  $n$ , the sequence of conditional error covariances  $\{\hat{P}_n(k)\}$  is stochastically bounded,*

$$\lim_{J \rightarrow \infty} \sup_{k \in \mathbb{T}_+} \mathbb{P}(\|\hat{P}_n(k)\| \geq J) = 0.$$

## Proof Methodology

- An appropriate stationary modification of the switched Riccati iterate governing the covariance evolution leads to a hypothetical Random Dynamical System (RDS) in the sense of Arnold.
- The RDS thus constructed is shown to be order preserving and strongly sublinear.
- Together with the global detectability and network connectivity assumptions, the weak convergence of the stationary hypothetical RDS is established.
- The ergodicity of the actual covariance sequence  $\{\hat{P}_n(k)\}$  is then applied to obtain the same convergence results for the original non-stationary system.

## Summary of Estimation Results

- Stability of distributed Kalman filtering errors is established under assumptions of network connectivity and global detectability.
- The analysis requires a new approach to the random Riccati equation (with non-stationary switching).
- Distributional convergence of the random Riccati equation is established;
- Approaching centralized performance exponentially fast over the inter-sensor communication rate.



## Distributed Detection

- $N$  sensors decide between two hypotheses:  $H_0$  vs.  $H_1$ .
- Sensor  $n$  observes  $Y_n(k)$  at time  $k$ ; observation vector:  $\mathbf{Y}(k) = [Y_1(k), \dots, Y_N(k)] \in \mathbb{R}^N$
- Conditional PDFs  $f_0$  and  $f_1$ :

$$\text{Under } H_l : f_l(\mathbf{Y}(k))d\nu, \quad l = 0, 1.$$

- Global Kullback-Lieber (KL) divergence  $D(P_0||P_1)$  vs.  $D(P_1||P_0)$  :

$$D(P_0||P_1) = \mathbb{E}_0 \left[ \ln \left( \frac{f_0(\mathbf{Y}(k))}{f_1(\mathbf{Y}(k))} \right) \right] = \int_{\mathbb{R}^N} \ln \left( \frac{f_0(\mathbf{Y})}{f_1(\mathbf{Y})} \right) f_0(\mathbf{Y})d\nu(\mathbf{Y}).$$

## Mixed Time Scale Distributed Detection: Algorithm $\mathcal{MD}$

- Each sensor maintains a local decision variable:  $X_n(k)$ .
- Connected sensor pair  $(i, j)$  communicates as  $U_{i,j}(k) = X_i(k) + W_{i,j}(k)$ .
- Local decision:  $H_1$  if  $X_n(k) > 0$ ; otherwise  $H_0$ .
- Type-I error:

$$\mathbb{P}_e^1(n, k) = \mathbb{P}_0(X_n(k) > 0),$$

and the Type-II error:

$$\mathbb{P}_e^2(n, k) = \mathbb{P}_1(X_n(k) \leq 0).$$

## Local Decision Variable $X_n(k)$

- At time  $k$ , local observation generates local intelligence (LLR):

$$S_n(k) = \log \left( \frac{f_1(Y_n(k))}{f_0(Y_n(k))} \right).$$

- At time  $k + 1$ , local decision variable is updated (mixed time scale) as

$$\begin{aligned} X_n(k+1) &= X_n(k) - \beta(k) \sum_{j \in \Omega_n} [X_n(k) - U_{n,j}(k)] + \alpha(k) [S_n(k) - X_n(k)] \\ &= X_n(k) - \beta(k) \sum_{j \in \Omega_n} [X_n(k) - X_j(k) - W_{n,j}(k)] \\ &\quad + \alpha(k) [S_n(k) - X_n(k)] \\ &= \left[ 1 - \beta(k) \sum_{j \in \Omega_n} 1 - \alpha(k) \right] X_n(k) + \beta(k) \sum_{j \in \Omega_n} X_j(k) \\ &\quad + \beta(k) \sum_{j \in \Omega_n} W_{n,j}(k) + \alpha(k) S_n(k). \end{aligned}$$

## Key Assumptions

- **(E.1)**: The sensor observations are conditionally independent, i.e.,

$$\text{Under } H_l : \quad f_l(\mathbf{Y}(k))d\nu = \prod_{n=1}^N f_{n,l}(Y_n(k))d\nu, \quad l = 0, 1,$$

- **(E.2)**: We assume global detectability, i.e.,  $D(P_0||P_1) > 0$  (and  $D(P_1||P_0) > 0$ ).
- **(E.3)**: The inter-sensor communication network is connected.
- **(E.4)**: The time varying weight sequences  $\{\beta(k)\}_{k=0}^{\infty}$  and  $\{\alpha(k)\}_{k=0}^{\infty}$  associated with the update process satisfy

$$\beta(k) = \frac{b_0}{(k+1)^\tau}, \quad \frac{1}{2} < \tau < 1, \quad \text{and} \quad \alpha(k) = \frac{a}{(k+1)}.$$

$(\lim_{k \rightarrow \infty} \beta(k)/\alpha(k) = \infty$ : Consensus potential vs. Innovation potential.)

## Main Theorem and Summary

**Theorem 2.** *Let assumptions (E.1)-(E.4) hold. Then,*

$$\lim_{k \rightarrow \infty} \mathbb{P}_e^1(n, k) = 0 \quad \text{and} \quad \lim_{k \rightarrow \infty} \mathbb{P}_e^2(n, k) = 0, \quad n = 1, \dots, N.$$

- When local observations are i.i.d., a single time scale distributed algorithm is achievable;
- With Gaussian noise assumptions, error decay exponent is obtainable.

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## About My School and Research Group

- Texas A&M University: One of the largest public schools in USA;
- Dwight Look Engineering School (USNews Rank 13th)
- Dept. of Electrical and Computer Engineering (USNews Rank 18th)
- My research Group:
  - 2 PostDoc, 9 Ph.D students, and 2 MS students
  - 8 NSF and DoD research projects
  - Main topics: network information theory, cognitive radio networks, and stochastic signal processing for smart grid
  - Openings: Postdocs and students for the coming year